

STOR 664 CWE 2020

Name: _____

ID#: _____

I pledge that I have neither given nor received unauthorized aid on this exam.

Signature: _____

- (1) Let x be the predictor representing the potassium/carbon atomic ratio (K/C, in %) and y be the response representing the amount of absorbed carbon monoxide (CO, in mole/mole C %). The data is shown in the following Table 1.

| <i>observation number</i> | x | y |
|---------------------------|------|------|
| 1 | 0.05 | 0.05 |
| 2 | 0.05 | 0.10 |
| 3 | 0.25 | 0.25 |
| 4 | 0.25 | 0.35 |
| 5 | 0.50 | 0.75 |
| 6 | 0.50 | 0.85 |
| 7 | 0.50 | 0.95 |
| 8 | 1.25 | 1.42 |
| 9 | 1.25 | 1.75 |
| 10 | 1.25 | 1.82 |
| 11 | 1.25 | 1.95 |
| 12 | 1.25 | 2.45 |
| 13 | 2.10 | 3.05 |
| 14 | 2.10 | 3.19 |
| 15 | 2.10 | 3.25 |
| 16 | 2.10 | 3.43 |
| 17 | 2.10 | 3.50 |
| 18 | 2.10 | 3.93 |
| 19 | 2.50 | 3.75 |
| 20 | 2.50 | 3.93 |
| 21 | 2.50 | 3.99 |
| 22 | 2.50 | 4.07 |

Based on the data table, consider the regression model $y_{ij} = x_i\theta_i + \epsilon_{ij}$ with $j = 1, \dots, n_i, i = 1, \dots, p$ and $n_1 + \dots + n_p = n$. The model can be expressed as, in a matrix form, $Y = X\theta + \epsilon$ where X is a $n \times p$ matrix whose entries are known constants, $\theta_1, \dots, \theta_p$ are unknown parameters, and the components ϵ_{ij} 's of ϵ are iid random variables with mean 0 and variance σ^2 (also an unknown parameter). Let

$$\bar{y}_{i\cdot} = \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij}, \quad \bar{y}_{\cdot\cdot} = \frac{1}{n} \sum_{i=1}^p n_i \bar{y}_{i\cdot}.$$

(1a) (6 points) Write down the form of X , and specify the numerical values of p and n_1, \dots, n_p .

(1b) (8 points) Express $(X^t X)^{-1}$ by their entries.

(1c) (6 points) Is this model a one-way ANOVA model? Explain why.

(1d) (8 points) Find the LSE $\hat{\theta}_i, i = 1, \dots, p$ numerically.

- (1e) (6 points) Let $\mu_i = x_i\theta_i, i = 1, \dots, p$ and assume ϵ_{ij} 's follow the common distribution $N(0, \sigma^2)$ [Note: This normality is assumed in Parts (1e) — (1i)]. Consider a F -test for $H_0: \mu_1 = \dots = \mu_p$ (reduced model 1) versus the full model H_1 . Complete the following ANOVA table by filling each empty cell with correct degree of freedom and corresponding mean square.

| <i>Source</i> | <i>Sum of Squares</i> | <i>D.F.</i> | <i>Mean Square</i> |
|----------------------|-----------------------|-------------|--------------------|
| <i>Full Model</i> | SSR_1 | | |
| <i>Reduced Model</i> | SSR_0 | | |
| <i>Difference</i> | $SSE_0 - SSE_1$ | | |
| <i>Residual</i> | SSE_1 | | |
| <i>Total</i> | $SSTO$ | | |

Table 2: ANOVA table for problem (1e) with two nested models

- (1f) (6 points) Following (1e), spell out the following sums of squares in terms of y_{ij} 's:

$$SSTO =$$

$$SSR_1 =$$

$$SSE_1 =$$

$$SSR_0 =$$

$$SSE_0 =$$

- (1g) (8 points) Under the reduced model 1, find the numerical values of LSE $\hat{\theta}_i, i = 1, \dots, p$.

- (1h) (8 point) Under the full model H_1 , write down SSE_1 numerically and use it to give the value of an unbiased estimate for σ^2 .

(1i) (8 points) To construct 95% simultaneous confidence intervals for θ_1 and θ_4 using Scheffé method, denoted by $M_1 \pm c_1 G$ and $M_4 \pm c_4 G$ respectively, where $G^2 = q F_{q, n-p; 1-\alpha}$, give the numerical values of midpoints M_1, M_4 , the coefficients c_1, c_4 , and specify q .

(1j) (8 points) Now let us consider the case of iid random errors ϵ_{ij} 's with mean 0 and variance σ^2 but **without** assuming the normal distribution for them. If we bootstrap the residuals e_{ij} 's under the full model H_1 , what are the numerical values for the *little population* in this case?

(1k) (6 points) Define $\beta_i = \frac{x_i \theta_i - x_{i-1} \theta_{i-1}}{x_i - x_{i-1}}$, $i = 2, 3, \dots, p$. Under the normality assumption for ϵ_{ij} 's given in (1e), consider testing $H_0^*: \beta_2 = \dots = \beta_p$ (reduced model 2) vs the full model H_1 , and complete the following ANOVA table by filling each empty cell with correct degree of freedom and corresponding mean square.

| <i>Source</i> | <i>Sum of Squares</i> | <i>D.F.</i> | <i>Mean Square</i> |
|----------------------|-----------------------|-------------|--------------------|
| <i>Full Model</i> | SSR_1 | | |
| <i>Reduced Model</i> | SSR_0 | | |
| <i>Difference</i> | $SSE_0 - SSE_1$ | | |
| <i>Residual</i> | SSE_1 | | |
| <i>Total</i> | $SSTO$ | | |

Table 3: ANOVA table for problem (1k) with two nested models

(1l) (8 points) Now let us consider the case of iid random errors ϵ_{ij} 's with mean 0 and variance σ^2 but **without** assuming the normal distribution for them. If we bootstrap the residuals e_{ij} 's under

the reduced model 2 H_0^* , what are the numerical values for the *little population* in this case?

(1m) (8 points) If we use bootstrap to test H_0^* in (1l) and reject H_0^* when $\frac{(SSE_0 - SSE_1)/q}{SSE_1/(n-p)} > r$. What is the value for r based on the given data? Describe the step-by-step bootstrap procedure for calculating the p -value represented by $P_{H_0^*} \left(\frac{(SSE_0 - SSE_1)/q}{SSE_1/(n-p)} > r \right)$.

(2) (6 points) Interpret the two reduced models 1 and 2 given in (1e) and (1k) respectively, i.e. what do they imply in the context for Problem (1)? Is one of them nested in another? If so, in what order?